Using Multiple Outcomes to Improve the Synthetic Control Method

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CMU

Joint work with Liyang Sun (UCL) and Avi Feller (UC Berkeley)

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 Re-weight control units ("synthetic control") to closely match treated unit's pre-treatment outcomes

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Often interested in effects on multiple outcomes

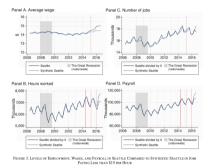
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- Incompatible SCs and potential over-fitting

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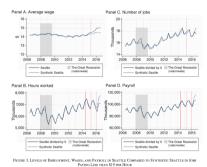
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- Fitting on all outcomes simultaneously
- Fitting on an index/avg of outcomes

Combines info across outcomes to reduce the bias



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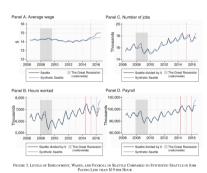
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Case study: Trejo et al. [2024] study on the 2014 Flint water crisis

- Math, reading, attendance, special needs



[Jardim et al., 2022]

Notation and estimands

Units:
$$i = 1, ..., N$$

Time:
$$t = 1, \dots, T$$

Outcomes:
$$k = 1, ..., K$$

$$k^{\text{th}}$$
 outcome for unit i at time t : Y_{itk}

First unit is treated at time T_0

Potential outcomes $Y_{itk}(0), Y_{itk}(1)$

$$\mathsf{treat} = \left(egin{array}{cccc} \checkmark & \checkmark & \checkmark \\ & & & \end{array} \right)$$

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Goal: Estimate effect on k^{th} outcome for treated unit at time $t \ge T_0$:

$$\tau_{tk} = Y_{1tk}(1) - Y_{1tk}(0)$$

$$\mathsf{treat} = \left(egin{array}{cccc} \checkmark & \checkmark & \checkmark \\ & & & \end{array} \right)$$

Synthetic control: weighted average of comparison units' outcomes

[Abadie et al., 2010, 2015]

$$\widehat{Y}_{1tk}(0) = \sum_{i \in \text{ctrls}} \hat{\gamma}_i Y_{itk}$$

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Weights optimize pre-treatment fit

$$\min_{\gamma \in \Delta} \sum_{t=1}^{T_0 - 1} \left(Y_{1tk} - \sum_{\text{controls}} \gamma_i Y_{itk} \right)^2$$

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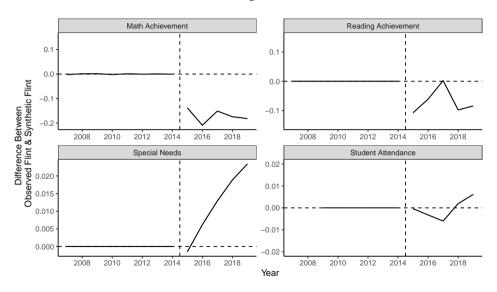
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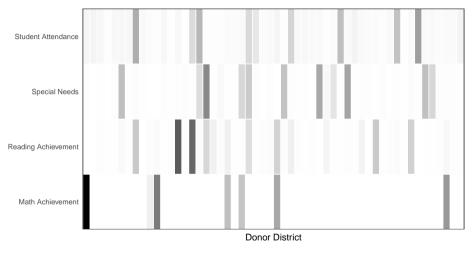
$$\min_{x \in A} \| \text{imbalance}_k \|_2^2$$

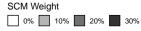
Abadie et al. [2010]: low bias if excellent pre-treatment fit and a long pre-period

Perfect fit on all outcomes → over-fitting?



Very different weights across outcomes \longrightarrow inconsistent analyses?





Typically assume a linear factor model: $Y_{it}(0) = \sum_{r=1}^{R} \phi_{ir} \mu_{tr} + \varepsilon_{it}$

- μ_t are J latent factors vary over time, fixed over units
- ϕ_i are J latent factor loadings vary over units, fixed over time

can't observe these

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Back to Flint: low # of time periods, might be overfitting + large approx error



Tackle both problems by using a common set of weights for outcomes $k = 1, \dots, K$

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Also include unit fixed effects (intercept-shifted SCM)

[Ferman and Pinto, 2021]

A shared latent structure across outcomes Link outcomes together via a common set of latent factor loadings

[in the paper: generalize this in terms of rank conditions]

$$Y_{itk}(0) = \alpha_{ik} + \beta_{tk} + \sum_{r=1}^{R} \phi_{ir} \mu_{tkr} + \epsilon_{itk}$$
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If R_0 common factors and ΔR idiosyncratic factors per outcome, sufficient condition:

$$R_0 + K \times \Delta R < N-1$$

- Test scores [Duflo et al., 2011]
- Finer temporal resolution [Sun, EBM & Feller (2024)]

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Gives a common set of **oracle** weights that balance the common latent factor loadings

$$\sum_{\text{controls}} \phi_i \gamma_i^* = \phi_{\text{trt}}$$

⁺ add'l regularity condition that $\|\gamma^*\|_1$ is bounded

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bias
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add'I bias from fixed effects like Nickell [1981] bias

- Approx error \downarrow as $T \uparrow$
- Pre-treatment fit stays the same

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For concatenated SCM weights:

$$bias < \frac{noise}{signal} + \frac{noise}{\sqrt{K}} \times \sqrt{\frac{\log N}{T}}$$

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For averaged SCM weights:

$$\mathsf{bias} < \frac{1}{\sqrt{K}} \times \left(\frac{ \underbrace{\mathsf{noise}}{\mathsf{signal}} + \mathsf{noise} \times \sqrt{\frac{\log N}{T}} \right)$$

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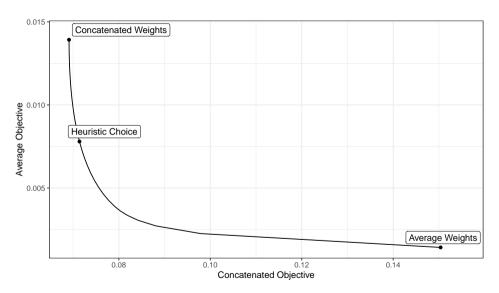
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Option 3: combined approach

$$\min_{\gamma \in \Delta} \nu \left\| \frac{1}{K} \sum_{k=1}^{K} \mathsf{imbalance}_{k} \right\|_{2}^{2} + \frac{1 - \nu}{K} \sum_{k=1}^{K} \left\| \mathsf{imbalance}_{k} \right\|_{2}^{2}$$

- In principle, a correct ν^* exists, but depends on the model
- Heuristic $\hat{\nu}$: ratio of avg and concatenated fit for concatenated SCM
- Vary u as a sensitivity parameter

The balance frontier



Adapt the conformal inference approach from Chernozhukov et al. [2021]

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Operates as a randomization test of a sharp null $H_0: (au_1, \dots, au_K) = (au_{10}, \dots, au_{K0})$

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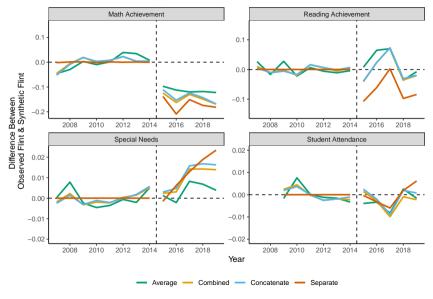
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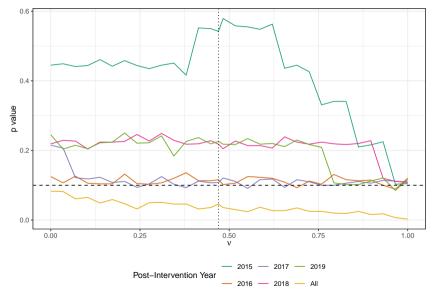
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- 3. Compute a test statistic on the residuals
- 4. Randomly scramble pre-post treatment time indicator and compute *p*-value by comparing observed test stat to the distribution
- Asymptotically correct size as $T \to \infty$
- But requires us to specify joint null on all outcomes together

Effects measured via different approaches



Sensitivity to ν



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- Common practice: run a separate SCM analysis for each outcome
- Practical and theoretical pitfalls: potential for overfitting, inconsistent analyses

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Many open questions and next steps

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- Weaken shared factor structure, e.g. hierarchical models?
- Less demanding form of inference?

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Thank you!

ebenmichael.github.io



References I

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